

2  $\delta' \gamma$

תורת המשפטים

(MME) ... תורת המשפטים

תורת המשפטים MME  $\theta$  (1)

$\theta = \sigma$ ,  $N(0, \sigma^2)$  (a)  $\theta = (m, p)$ ,  $B(m, p)$  (a)

$\theta = (\mu, \sigma^2)$ ,  $N(\mu, \sigma^2)$  (d)  $\theta = (\alpha, \lambda)$ ,  $\Gamma(\alpha, \lambda)$  (c)

$\theta = b - a$ ,  $U(a, b)$  (f)  $\theta = (a, b)$ ,  $U(a, b)$  (e)

$\theta = a$ ,  $U(-a, a)$  (h)  $\theta = a$ ,  $U(a, 2a)$  (g)

$\theta = (m, p)$ ,  $NB(m, p)$  (j)  $\theta = m$ ,  $\chi_m^2 = \Gamma(\frac{m}{2}, \frac{1}{2})$  (i)

תורת המשפטים MME  $\theta$  (2)

$f(x, \theta) = \frac{1}{\theta} - \frac{|x|}{\theta^2}$ ,  $|x| \leq \theta$  (a)

$\theta = (\mu, \sigma^2)$ , lognormal  $(\mu, \sigma^2)$  (b)

Beta  $(\theta, 1)$ ,  $f(x, \theta) = \theta x^{\theta-1}$ ,  $0 < x \leq 1$  (c)

$f(x, \theta) = \frac{2(\theta-x)}{\theta^2}$ ,  $0 \leq x \leq \theta$  (d)

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$M'_r = \frac{1}{n} (X_1^r + X_2^r + \dots + X_n^r)$

$M'_1 = \bar{X}$ ,  $r = 1, 2, 3, \dots$

Table 1 DISCRETE DISTRIBUTIONS

Name of parametric family of distributions	Discrete density functions $f(x)$	Parameter space	Mean $\mu = E(X)$
Discrete uniform $U_d(N)$	$f(x) = \frac{1}{N} I_{\{1, \dots, N\}}(x)$	$N = 1, 2, \dots$	$\frac{N+1}{2}$
Bernoulli $B(1, p)$	$f(x) = p^x q^{1-x} I_{\{0, 1\}}(x)$	$0 \leq p \leq 1$ $(q = 1 - p)$	$p$
Binomial $B(n, p)$	$f(x) = \binom{n}{x} p^x q^{n-x} I_{\{0, 1, \dots, n\}}(x)$	$0 \leq p \leq 1$ $n = 1, 2, 3, \dots$ $(q = 1 - p)$	$np$
Hypergeometric $H(M, K, n)$	$f(x) = \frac{\binom{K}{x} \binom{M-K}{n-x}}{\binom{M}{n}} I_{\{0, 1, \dots, n\}}(x)$	$M = 1, 2, \dots$ $K = 0, 1, \dots, M$ $n = 1, 2, \dots, M$	$\frac{K}{M}$
Poisson $P(\lambda)$	$f(x) = \frac{e^{-\lambda} \lambda^x}{x!} I_{\{0, 1, \dots\}}(x)$	$\lambda > 0$	$\lambda$
Geometric $G(p)$	$f(x) = p q^{x-1} I_{\{1, 2, \dots\}}(x)$	$0 < p \leq 1$ $(q = 1 - p)$	$\frac{1}{p}$
Negative binomial $NB(m, p)$	$f(x) = \binom{x-1}{m-1} p^m q^{x-m} I_{\{m, m+1, \dots\}}(x)$	$0 < p \leq 1$	$\frac{m}{p}$

$$VS^2 = \frac{1}{n} \left( \mu_4 - \frac{n-3}{n-1} \sigma^4 \right)$$

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = \frac{1}{n-1} \sum_{i=1}^n X_i^2 - \frac{n}{n-1} \bar{X}^2$$

Variance $\sigma^2 = E[(X - \mu)^2]$	Moments $\mu_r^* = E[(X^r)]$ or $\mu_r = E[(X - \mu)^r]$ and/or cumulants $\kappa_r$	Moment generating function $e^{t(X-\mu)}$
$\frac{N^2 - 1}{12}$	$\mu_1^* = \frac{N(N+1)^2}{4}$ $\mu_2^* = \frac{(N+1)(2N+1)(3N+1)}{30}$	$\sum_{j=1}^N \frac{1}{N} e^{jt}$
$npq$	$\mu_1^* = p$ for all $r$ $\mu_3 = npq(q-p)$ $\mu_4 = 3n^2 p^2 q^2 + npq(1-6pq)$	$q + pe^t$ $(q + pe^t)^n$
$\frac{KM - K}{M} \frac{M - n}{M - 1}$	$E[X(X-1) \dots (X-r+1)] = r! \frac{\binom{K}{r} \binom{M}{n-r}}{\binom{M}{n}}$	not useful
$\lambda$	$\mu_r = \lambda$ for $r = 1, 2, \dots$ $\mu_3 = \lambda$ $\mu_4 = \lambda + 3\lambda^2$	$\exp[\lambda(e^t - 1)]$
$\frac{q}{p^2}$	$\mu_1 = \frac{q + q^2}{p^2}$ $\mu_2 = \frac{q + 7q^2 + q^3}{p^3}$	$\frac{pe^t}{1 - qe^t}$
$\frac{q^2}{p^3}$	$\mu_3 = \frac{q(1+q^2)}{p^3}$ $\mu_4 = \frac{q(q + (3r+4)q^2 + q^3)}{p^4}$	

$$B(a, \theta) = \frac{\Gamma(a) \cdot \Gamma(\theta)}{\Gamma(a+\theta)}$$

$$\Gamma(n) = (n-1)!$$

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Table 2 CONTINUOUS DISTRIBUTIONS

Name of parametric family of distributions	Cumulative distribution function $F(x)$ or probability density function $f(x)$	Parameter space	Mean $\mu = E\{X\}$
Uniform or rectangular $U(a, b)$	$f(x) = \frac{1}{b-a} I_{(a,b)}(x)$	$-\infty < a < b < \infty$	$\frac{a+b}{2}$
Normal $N(\mu, \sigma^2)$	$f(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp[-(x-\mu)^2/2\sigma^2]$	$-\infty < \mu < \infty$ $\sigma > 0$	$\mu$
Exponential $EXP(\lambda)$	$f(x) = \lambda e^{-\lambda x} I_{(0,\infty)}(x)$	$\lambda > 0$	$\frac{1}{\lambda}$
Gamma $\Gamma(r, \lambda)$	$f(x) = \frac{\lambda^r}{\Gamma(r)} x^{r-1} e^{-\lambda x} I_{(0,\infty)}(x)$	$\lambda > 0$ $r > 0$	$\frac{r}{\lambda}$
Beta $Beta(a, b)$	$f(x) = \frac{1}{B(a, b)} x^{a-1} (1-x)^{b-1} I_{(0,1)}(x)$	$a > 0$ $b > 0$	$\frac{a}{a+b}$
Cauchy	$f(x) = \frac{1}{\pi\beta} \frac{1}{1 + [(x-a)/\beta]^2}$	$-\infty < a < \infty$ $\beta > 0$	Does not exist
Lognormal	$f(x) = \frac{1}{x\sqrt{2\pi}\sigma} \exp[-(\log x - \mu)^2/2\sigma^2] I_{(0,\infty)}(x)$	$-\infty < \mu < \infty$ $\sigma > 0$	$\exp\{\mu + \frac{1}{2}\sigma^2\}$
Double exponential	$f(x) = \frac{1}{2\beta} \exp\left(-\frac{ x-a }{\beta}\right)$	$-\infty < a < \infty$ $\beta > 0$	$a$

$$B(a, b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$

$$\int_0^1 x^{\alpha-1} (1-x)^{\beta-1} dx$$

Variance $\sigma^2 = E\{(X-\mu)^2\}$	Moments $\mu_r^r = E\{X^r\}$ or $\mu_r = E\{(X-\mu)^r\}$ and/or cumulants $\kappa_r$	Moment generating function $E\{e^{tX}\}$
$\frac{(b-a)^2}{12}$	$\mu_r = 0$ for $r$ odd $\mu_r = \frac{(b-a)^r}{2(r+1)}$ for $r$ even	$\frac{e^{bt} - e^{at}}{(b-a)t}$
$\sigma^2$	$\mu_r = 0, r$ odd; $\mu_r = \frac{r!}{(r/2)! 2^{r/2}} \sigma^2, r$ even; $\kappa_r = 0, r > 2$	$\exp\{\mu t + \frac{1}{2}\sigma^2 t^2\}$
$\frac{1}{\lambda^2}$	$\mu_r^r = \frac{\Gamma(r+1)}{\lambda^r}$	$\frac{\lambda}{\lambda-t}$ for $t < \lambda$
$\frac{c}{\lambda^2}$	$\mu_3 = \frac{\Gamma(r+1)}{\lambda^3 \Gamma(r)} = \frac{3}{\lambda} \frac{\Gamma(r) \Gamma(r-1)}{\Gamma(r)}$ for $r < \lambda$ $\mu_r = \frac{r!}{\lambda^r} B(r+a, b)$	not useful
Does not exist	Does not exist	Characteristic function is $e^{\mu t - \beta t }$
$\frac{ab}{(a+b+1)(a+b)^2}$	Does not exist	not useful
Does not exist	Does not exist	not useful
$2\beta^2$	$\mu_r = 0$ for $r$ odd; $\mu_r = r! \beta^r$ for $r$ even	$\frac{e^{-t}}{1-(\beta t)^2}$

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Table 2. CONTINUOUS DISTRIBUTIONS (continued)

Name of parametric family of distributions	Cumulative distribution function $F(x)$ or probability density function $f(x)$	Parameter space	Mean $\mu = E(X)$
Weibull $W(c, \theta)$	$f(x) = \theta x^{\theta-1} \exp(-ax^\theta) I_{(0, \infty)}(x)$	$a > 0$ $\theta > 0$	$\mu = \frac{\Gamma(1 + \frac{1}{\theta})}{\theta}$
Logistic	$F(x) = [1 + e^{-(x-\mu)/\beta}]^{-1}$	$-\infty < x < \infty$ $\beta > 0$	$\alpha$
Pareto $Par(\theta > x_0)$	$f(x) = \frac{\theta x_0^\theta}{x^{\theta+1}} I_{(x_0, \infty)}(x)$	$x_0 > 0$ $\theta > 0$	$\frac{\theta x_0}{\theta - 1}$ (for $\theta > 1$ )
Clumbel or extreme value	$F(x) = \exp(-e^{-(x-\mu)/\beta})$	$-\infty < x < \infty$ $\beta > 0$	$\alpha + \beta \gamma$ $\gamma \approx .577216$
t distribution	$f(x) = \frac{\Gamma(k + 1/2)}{\Gamma(k/2)} \frac{1}{\sqrt{k\pi}} \frac{1}{(1 + x^2/k)^{k+1/2}}$	$k > 0$	$\mu = 0$ (for $k > 1$ )
F distribution	$f(x) = \frac{\Gamma(m+n/2) \Gamma(m/n)}{\Gamma(m/2) \Gamma(n/2)} \frac{x^{m-1}}{[1 + (m/n)x^2]^{(m+n)/2}} I_{(0, \infty)}(x)$	$m, n = 1, 2, \dots$	$\frac{n}{n-2}$ (for $n > 2$ )
Chi-square distribution	$f(x) = \frac{1}{\Gamma(k/2)} \left(\frac{x}{2}\right)^{k/2-1} e^{-(x/2)} I_{(0, \infty)}(x)$	$k = 1, 2, \dots$	$k$

Variance $\sigma^2 = E[(X - \mu)^2]$	Moments $\mu_r = E[X^r]$ or $\mu_r = E[(X - \mu)^r]$ and/or cumulants $\kappa_r$	Moment generating function $E[e^{tX}]$
$\frac{a^{-2n} \Gamma(n + 2b^{-1})}{\Gamma(n + b^{-1})}$	$\mu_r = a^{-rn} \Gamma\left(1 + \frac{r}{b}\right)$	$E[e^{tX}] = a^{-rt} \Gamma\left(1 + \frac{t}{b}\right)$
$\frac{\beta^2 \pi^2}{3}$		$e^{t\mu} \beta t \operatorname{csc}(\pi\beta t)$
$\frac{\theta x_0^2}{(b-1)^2(b-2)}$ (for $b > 2$ )	$\mu_r = \frac{\theta x_0^r}{b-r}$ for $\theta > r$	does not exist
$\frac{\pi^2 \beta^2}{6}$	$\kappa_2 = (-\beta)^2 \psi^{(1)}(1)$ for $r \geq 2$ , where $\psi(x)$ is digamma function	$E[e^{tX}] = a^{-rt} \Gamma(1 - \beta t)$ for $t < 1/\beta$
$\frac{k}{k-2}$ (for $k > 2$ )	$\mu_2 = 0$ for $k > r$ and $r$ odd $\mu_2 = \frac{k\pi^2 B(r + 1/2, (k-r)/2)}{B(k/2, k/2)}$ for $k > r$ and $r$ even	does not exist
$\frac{2n^2(n+n-2)}{n(n-2)(n-4)}$ (for $n > 4$ )	$\mu_4 = \left(\frac{n}{n-4}\right) \frac{\Gamma(n/2 + r) \Gamma(n/2 - r)}{\Gamma(n/2) \Gamma(n/2)}$ for $r < 2$	does not exist
$2k$	$\mu_2 = \frac{2\Gamma(k/2 + D)}{\Gamma(k/2)}$	$\left(\frac{1}{1-2t}\right)^{k/2}$ for $t < 1/2$

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$$\theta = (m, p)$$

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$$B(m, p)$$

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(5)

$$\mu_1' = mp$$

$$\mu_2' = mpq + (mp)^2 = \mu_1'q + (\mu_1')^2$$

$$\mu_1'q = \mu_2' - (\mu_1')^2$$

$$p = 1 - q = 1 - \frac{\mu_2' - (\mu_1')^2}{\mu_1'} = \frac{\mu_1' - \mu_2' + (\mu_1')^2}{\mu_1'}$$

$$m = \frac{\mu_1'}{p} = \frac{(\mu_1')^2}{\mu_1' - \mu_2' + (\mu_1')^2}$$

MME  
 $\bar{X} = \frac{1}{n} \sum_{k=1}^n X_k = M_1'$   
 $M_2' = \frac{1}{n} \sum_{k=1}^n X_k^2$

$$\hat{p} = \frac{\bar{X} - M_2' + (\bar{X})^2}{\bar{X}}, \quad \hat{m} = \frac{(\bar{X})^2}{\bar{X} - M_2' + (\bar{X})^2}$$

$$\hat{\sigma}^2 = M_2' - (\bar{X})^2, \quad \sigma = \theta, \quad \mathcal{N}(0, \sigma^2) \quad (6) \\ \hat{\sigma} = \sqrt{M_2' - (\bar{X})^2}$$

$$\mu_1' = \frac{\alpha}{\lambda}, \quad \mu_2' = \frac{\alpha(\alpha+1)}{\lambda^2}, \quad \theta = (\alpha, \lambda), \quad \Gamma(\alpha, \lambda) \quad (6) \\ ///$$

$$\frac{\mu_2'}{\mu_1'} = \frac{\alpha+1}{\lambda} = \mu_1' + \frac{1}{\lambda}, \quad \frac{1}{\lambda} = \frac{\mu_2' - (\mu_1')^2}{\mu_1'}$$

$$\lambda = \frac{\mu_1'}{\mu_2' - (\mu_1')^2}, \quad \alpha = \lambda \mu_1' = \frac{(\mu_1')^2}{\mu_2' - (\mu_1')^2}$$

$$\hat{\lambda} = \frac{\bar{X}}{M_2' - (\bar{X})^2}, \quad \hat{\alpha} = \frac{(\bar{X})^2}{M_2' - (\bar{X})^2}$$

$$\hat{\mu} = \bar{X}, \quad \theta = (\mu, \sigma^2) \quad \mathcal{N}(\mu, \sigma^2) \quad (6) \\ \hat{\sigma}^2 = M_2' - (\bar{X})^2$$

$$\theta = (a, b), \quad U(a, b)$$

(e) (1) (2827) (6)

$$\mu_1' = \frac{a+b}{2}$$

$$\mu_2' = \frac{a^2 + ab + b^2}{3}$$

$$\hat{a} = \bar{X} - \sqrt{3} \sqrt{\mu_2' - (\bar{X})^2}$$

$$\hat{b} = \bar{X} + \sqrt{3} \sqrt{\mu_2' - (\bar{X})^2}$$

$$\theta = b - a, \quad U(a, b) \quad (f)$$

$$\hat{\theta} = 2\sqrt{3} \sqrt{\mu_2' - (\bar{X})^2}$$

$$\hat{\theta} = \frac{2}{3} \bar{X}$$

$$\theta = a, \quad U(a, 2a) \quad (g)$$

$$\mu_1' = \frac{3}{2} a$$

$$\mu_1' = 0$$

$$\mu_2' = \frac{(2a)^2}{12} = \frac{a^2}{3}$$

$$\theta = a, \quad U(-a, a) \quad (h)$$

$$\hat{a} = \sqrt{3 \mu_2'}$$

$$\theta = n, \quad \chi_n^2 \sim \Gamma\left(\frac{n}{2}, \frac{1}{2}\right) \quad (i)$$

$$\mu_2' = n, \quad \hat{n} = \bar{X}$$

$$NB(m, p) \quad (j)$$

$$\theta = (m, p)$$

$$\mu_1' = \frac{m}{p}$$

$$\mu_2' = \frac{mq}{p^2} + \left(\frac{m}{p}\right)^2 = \frac{m(q+m)}{p^2}$$

$$\mu_2' = \mu_1' \left(\frac{m+q}{p}\right) = \mu_1' \left(\mu_1' + \frac{q}{p}\right)$$

$$\frac{q}{p} = \frac{\mu_2'}{\mu_1'} - \mu_1', \quad m = p \mu_1' = \frac{(\mu_1')^2}{\mu_2' - (\mu_1')^2 + \mu_1'}$$

$$p = \frac{\mu_1'}{\mu_2' - (\mu_1')^2 + \mu_1'}$$

$$\hat{p} = \frac{-\bar{X}}{\mu_2' - \bar{X}^2 + \bar{X}}$$

$$\hat{m} = \frac{\bar{X}^2}{\mu_2' - \bar{X}^2 + \bar{X}}$$

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$0 < \theta, |x| \leq \theta \quad f(x, \theta) = \frac{\theta - |x|}{\theta^2}$

(a) (2)

$M_1' = 0 \quad M_2' = 2 \int_0^\theta x^2 \frac{\theta - x}{\theta^2} dx = \frac{2}{\theta^2} \left( \frac{x^3 \theta}{3} - \frac{x^4}{4} \right) \Big|_0^\theta =$

$M_2' = \frac{\theta^2}{6} \quad \hat{\theta} = \sqrt{6 M_2'} = \sqrt{\frac{6}{n} \sum_{k=1}^n X_{k0}^2}$

$\theta = (\mu, \sigma^2)$

lognormal  $(\mu, \sigma^2)$  (b)

$M_1' = e^{\mu + \frac{1}{2} \sigma^2}$

$2\mu + \sigma^2 = 2 \ln M_1'$

$M_2' = e^{2\mu + 2\sigma^2}$

$2\mu + 2\sigma^2 = \ln M_2'$

$\mu = \frac{\ln M_2' - \sigma^2}{2}$

$\sigma^2 = \ln M_2' - 2 \ln M_1'$

$\sigma^2 = \ln \frac{M_2'}{(M_1')^2}$

$\mu = \frac{2 \ln M_1' - \ln M_2'}{2}$

$\hat{\sigma}^2 = \ln \frac{M_2'}{(\bar{X})^2}$

$\hat{\mu} = \frac{1}{2} \ln \frac{\bar{X}^4}{M_2'}$

$M_1' = \frac{\theta}{\theta + 1}$

$0 \leq x \leq L, f(x, \theta) = \theta x^{\theta-1}$

(c)

$\theta = \frac{M_1'}{1 - M_1'}$

$\hat{\theta} = \frac{\bar{X}}{1 - \bar{X}} \quad \text{Beta}(\theta, 1)$

$0 \leq x \leq \theta, f(x, \theta) = \frac{2}{\theta^2} (\theta - x)$

(d)

$M_1' = E X = \frac{2}{\theta^2} \int_0^\theta x(\theta - x) dx = \frac{\theta}{3}$

$\hat{\theta} = 3 \bar{X}$